# Some New Tables for Upper Probability Points of the Largest Root of a Determinantal Equation with Seven and Eight Roots

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We revisit the Fisher-Girshick-Hsu-Roy distribution (1939), which has interested statisticians for more than six decades. Instead of using K.C.S. Pillai's method of neglecting higher order terms of the cumulative distribution function (C.D.F.) of the largest root to approximate the percentage points, we simply keep the whole C.D.F. and apply its natural non-decreasing property to calculate the exact probabilities. At the duplicated percentage points, we found our computed percentage points to be consistent with existing tables. However, our tabulations have greatly extended the existing tables.

In Chen (2002), we were concerned with the distribution of the largest characteristic roots in multivariate analysis when there are two to six roots. Now, we will extend the size to seven and eight roots.Fisher-Girshick-Hsu-Roy(1939) discuss this in detail and present the joint probability density function in general. This well-known distribution depends on the number of characteristic roots and two parameters m and n, which are defined differently for various situations, as described by Pillai (1955). The upper percentage points of the distribution are commonly used in three different multivariate hypothesis tests: tests of equality of the variance-covariance matrices of two p-variate normal populations, tests of equality of the p-dimensional mean vectors for k p-variate normal populations, and tests of independence between a p-set and a q-set of variates in a (p+q)-variate normal population. When the null hypotheses are true, these three proposed tests depend only on the characteristic roots of matrices using observed samples. The problem can be stated as follows: using a random sample from the multivariate normal population, we will compute the characteristic roots from a sum of product matrices of this sample. We will then compare the largest characteristic root of the matrices with the percentage points tabulated in this paper to determine whether or not the null hypothesis is rejected at a certain probability confidence.

There are already many published tables that focus on upper percentage point tabulations or chart the various sizes of roots. The best-known contributor in this area is Pillai, who gave general rules for finding the C.D.F. of the largest root and tabulated upper percentage points of 95% and 99% for various sizes of roots. Other contributors, including Nanda (1948, 1951), Foster and Rees (1957, 1958), and Heck (1960) will be discussed in more detail in section 2. Section 3 contains the joint distribution of s non-null characteristic roots of a matrix in general form and the C.D.F. of the seven and eight largest characteristic roots. The algorithm used to create the tables in this paper is the same as in Chen (2003), and we will not repeat it. Also, we will ignore the discussion of precision of the results.

## ▶ Cumulative Function and Historical Work

The joint distribution of s non-null characteristic roots of a matrix in multivariate distribution was first given by Fisher-Girshick-Hsu-Roy (1939) and can be expressed in the form of (3.1). We further extended the distribution of the largest characteristic root to seven and eight roots. Even though the form of the joint density function is known, it is not easy to write out the C.D.F. of the largest characteristic root to seven roots. To solve this problem, two methods can be used to find the C.D.F. more easily. Pillai (1965) suggests that the C.D.F. of the largest characteristic root could be presented in determinant form of incomplete beta functions. Since the numerical integration of each of the s factorial multiple integrals when the determinant is expanded is difficult, he suggests an alternative reduction formula that gives exact expressions for the C.D.F. of the largest root in terms of incomplete beta functions or functions of incomplete beta functions for various values of s. An alternative method suggested by Nanda (1948) yields the same results. He started with the Vandermonde determinant and expanded it in minors of a row, then repeated applied integration by part to find the C.D.F. of the largest characteristic root. In this paper, we use the Pillai notation and present the case with seven roots in equation (3.2). Following this C.D.F. and the algorithm previously used, we tabulate the upper percentage points.

Here, it is useful to review some of the published tables and reasons to extend the tables. Pillai (1956a, 1959) published tables that focus only on two percentage points: 95% and 99% for s =2,6, m = 0(1)4, and n varying from 5 to 1000. Foster and Rees (1957) tabulated the upper percentage points 80%, 85%, 90%, 95%, and 99% of the largest root for s=2, m=-0.5, 0(1)9, n=1(1)19 (5)49,59,79. Foster(1957, 1958) further extended these tables for values of s=3 and 4. Heck (1960) has given some charts of upper 95%, 97.5%, and 99% points for s=2(1)5, m=-0.5, 0(1)10, and n greater than 4. These table values can be applied to our statistical analysis with some standard textbooks as references. For example, recently, Rencher included the percentage point 0.950 in two textbooks Rencher (1998 and 2002).

Without a modern computer, it is difficult and tedious to compute the whole C.D.F. (3.2) at each percentage point. Therefore, deleting higher order terms and retaining a few lower order terms to approximate the roots is a reasonable solution. However, this approach involves intolerable error at lower percentage points, such as 80%, 82.5%, 85%, 87.5%, 90%, or 92.5%. These percentage points are usually ignored due to the difficulty of their computation, and not due to their lack of use. Traditional methods treat intermediate percentage points by interpolation, but without, for example, 85% or 90% percentage points, it is difficult to interpolate 87.5%. In recent years, computers have gradually matured in memory, speed, and flexibility in usage, which has greatly changed the methods by which we study statistics. In this study, we use one of the most basic properties of C.D.F. and revisit this most important distribution. As many percentage points as are needed in one computer run are included: these are 0.80, 0.825, 0.850, 0.875, 0.890, 0.900, 0.910(0.005), 0.995. Different authors have selected different m and n parameter values, but we selected these parameters such that all existing table values are included. For the parameters m=0(1)10 and n=3(1)20(2)30(5)80(10), 150, 200(100)1000, our table provides the percentage points and probabilities while avoiding the interpolation problem.

## ▶ The Distribution Function of Seven Characteristic Roots

Suppose $x = \{x_{ij}\}$ and $x^* = \{x^*_{ij}\}$ are two p-variate random matrices with $n_1$ and $n_2$ the degree of freedom, respectively. Assume the two multivariate populations have the same covariance matrix: for example, $S_1 = xx^T / n_1$ and $S_2 = x^* x^{*T} / n_2$. When the null hypothesis is true, both $S_1$ and $S_2$ are independent estimators of the unknown but equal covariance matrices. The joint distribution of the roots of the determinantal equation $|A - \theta(A+B)| = 0$ where $A = n_1 S_1$ and $B = n_2 S_2$ has been given by Fisher-Girshick-Hsu-Roy(1939) and can be written as :

$$f(\theta_1, \ldots \theta_s) = C(s,m,n) \prod_{i=1}^{s} \theta^m_i (1-\theta_i)^n \prod_{i>j} (\theta_i - \theta_j)$$

$$(0 < \theta_1 \leq \ldots \leq \theta_s \leq 1),$$

where

$$C(s,m,n) = \frac{\pi^{s/2} \prod_{i=1}^{s} \Gamma\left(\frac{2m+2n+s+i+2}{2}\right)}{\prod_{i=1}^{s} \Gamma\left(\frac{2m+i+1}{2}\right) \Gamma\left(\frac{2n+i+1}{2}\right) \Gamma\left(\frac{i}{2}\right)}. \quad (3.1)$$

and the parameters m and n are defined differently for various situations as described by Pillai (1955, pp. 118). Following Pillai's method, the cumulative distribution function of the largest characteristic root for seven and eight is given below:

When s = 7, the C.D.F. of the largest characteristic root is:

$$\begin{aligned} \Pr(\theta_7 \leq x) = \frac{C(7,m,n)}{m+n+7} [ & -IO(x,m+6,n+1) * v\_0654321x(x,m,n) \\ & -2I(x,2m+6,2n+1) * v\_054321x(x,m+1,n) + 2I(x,2m+7,2n+1) \\ & * v\_05432x(x,m,n) - 2I(x,2m+8,2n+1) * v\_05431x(x,m,n) \\ & + 2I(x,2m+9,2n+1) * v\_05421x(x,m,n) - 2I(x,2m+10,2n+1) \\ & * v\_05321x(x,m,n) + 2I(x,2m+11,2n+1) * v\_054321x(x,m,n)] \end{aligned}$$

$$\Pr(\theta_8 \le x) = \frac{C(8,m,n)}{m+n+8}[-IO(x,m+7,n+1)*v\_07654321x(x,m,n)$$
$$+2I(x,2m+7,2n+1)*v\_0654321x(x,m+1,n)-2I(x,2m+8,2n+1)$$
$$*v\_065432x(x,m,n)+2I(x,2m+9,2n+1)*v\_065431x(x,m,n)$$
$$-2I(x,2m+10,2n+1)*v\_065421x(x,m,n)+2I(x,2m+11,2n+1)$$
$$*v\_065321x(x,m,n)-2I(x,2m+12,2n+1)*v\_064321x(x,m,n)$$
$$+2I(x,2m+13,2n+1)*v\_0654321x(x,m,n)]$$

$$C(7,m,n)=\frac{\Gamma(2m+2n+14)*\Gamma(2m+2n+12)*\Gamma(2m+2n+10)*\Gamma(2m+2n+8)}{46080*\Gamma(2m+6)*\Gamma(2n+6)*\Gamma(2m+4)*\Gamma(2n+4)*\Gamma(2m+2)*\Gamma(2n+2)*\Gamma(m+4)*\Gamma(n+4)}$$

$$C(8,m,n)=\frac{\Gamma(2m+2n+17)*\Gamma(2m+2n+15)*\Gamma(2m+2n+13)*\Gamma(2m+2n+11)}{8847360*\Gamma(2m+8)*\Gamma(2n+8)*\Gamma(2m+6)*\Gamma(2n+6)*\Gamma(2m+4)*\Gamma(2n+4)*\Gamma(2m+2)*\Gamma(2n+2)}$$

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**Upper percentage points of .900 of theta(p,m,n), the largest eigenvalue of |B-theta(W+B)|=0, when s=7**

| | m | | | | | | | | | | |
|---|---|---|---|---|---|---|---|---|---|---|---|
| n | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 |
| 3 | .9040 | .9188 | .9295 | .9378 | .9442 | .9495 | .9538 | .9576 | .9608 | .9630 | .9644 |
| 4 | .8650 | .8842 | .8986 | .9097 | .9186 | .9259 | .9320 | .9371 | .9415 | .9453 | .9490 |
| 5 | .8266 | .8497 | .8671 | .8809 | .8920 | .9012 | .9090 | .9156 | .9212 | .9261 | .9307 |
| 6 | .7899 | .8160 | .8362 | .8522 | .8653 | .8763 | .8855 | .8935 | .9004 | .9064 | .9116 |
| 7 | .7552 | .7838 | .8062 | .8242 | .8391 | .8515 | .8622 | .8714 | .8794 | .8865 | .8927 |
| 8 | .7226 | .7533 | .7774 | .7971 | .8135 | .8273 | .8392 | .8495 | .8586 | .8665 | .8737 |
| 9 | .6923 | .7244 | .7501 | .7711 | .7888 | .8038 | .8168 | .8281 | .8380 | .8469 | .8548 |
| 10 | .6639 | .6973 | .7241 | .7463 | .7650 | .7811 | .7950 | .8072 | .8180 | .8276 | .8363 |
| 11 | .6376 | .6717 | .6995 | .7226 | .7423 | .7592 | .7740 | .7869 | .7985 | .8088 | .8181 |
| 12 | .6130 | .6478 | .6763 | .7002 | .7206 | .7382 | .7537 | .7674 | .7796 | .7905 | .8004 |
| 13 | .5901 | .6253 | .6543 | .6788 | .6999 | .7182 | .7342 | .7485 | .7613 | .7727 | .7832 |
| 14 | .5687 | .6042 | .6336 | .6586 | .6801 | .6989 | .7155 | .7303 | .7436 | .7556 | .7664 |
| 15 | .5487 | .5843 | .6140 | .6394 | .6613 | .6806 | .6976 | .7128 | .7266 | .7390 | .7503 |
| 16 | .5300 | .5656 | .5955 | .6211 | .6434 | .6630 | .6804 | .6960 | .7101 | .7229 | .7346 |
| 17 | .5124 | .5480 | .5780 | .6038 | .6263 | .6462 | .6640 | .6799 | .6943 | .7074 | .7194 |
| 18 | .4959 | .5314 | .5614 | .5873 | .6100 | .6302 | .6482 | .6644 | .6791 | .6925 | .7048 |
| 19 | .4805 | .5157 | .5457 | .5717 | .5945 | .6148 | .6330 | .6495 | .6644 | .6781 | .6906 |
| 20 | .4659 | .5009 | .5308 | .5568 | .5797 | .6001 | .6185 | .6351 | .6503 | .6642 | .6769 |
| 22 | .4391 | .4736 | .5032 | .5291 | .5520 | .5726 | .5912 | .6081 | .6236 | .6378 | .6509 |
| 24 | .4152 | .4490 | .4782 | .5039 | .5268 | .5474 | .5661 | .5832 | .5988 | .6133 | .6267 |
| 26 | .3937 | .4267 | .4554 | .4809 | .5036 | .5242 | .5429 | .5600 | .5758 | .5904 | .6040 |
| 28 | .3743 | .4065 | .4347 | .4598 | .4823 | .5027 | .5214 | .5386 | .5544 | .5691 | .5828 |
| 30 | .3567 | .3881 | .4158 | .4404 | .4627 | .4829 | .5015 | .5186 | .5344 | .5492 | .5629 |
| 35 | .3190 | .3486 | .3748 | .3984 | .4198 | .4394 | .4576 | .4744 | .4901 | .5047 | .5184 |
| 40 | .2885 | .3162 | .3410 | .3635 | .3840 | .4030 | .4205 | .4369 | .4523 | .4667 | .4802 |
| 45 | .2632 | .2894 | .3128 | .3342 | .3538 | .3720 | .3889 | .4048 | .4197 | .4338 | .4471 |
| 50 | .2420 | .2666 | .2888 | .3092 | .3279 | .3454 | .3617 | .3770 | .3915 | .4052 | .4181 |
| 55 | .2240 | .2472 | .2683 | .2876 | .3055 | .3223 | .3380 | .3528 | .3668 | .3800 | .3926 |
| 60 | .2084 | .2304 | .2504 | .2688 | .2860 | .3020 | .3171 | .3314 | .3449 | .3578 | .3700 |
| 65 | .1949 | .2157 | .2348 | .2524 | .2688 | .2842 | .2987 | .3124 | .3255 | .3379 | .3498 |
| 70 | .1830 | .2028 | .2210 | .2378 | .2535 | .2683 | .2822 | .2955 | .3081 | .3202 | .3317 |
| 75 | .1725 | .1914 | .2087 | .2248 | .2399 | .2541 | .2675 | .2803 | .2925 | .3042 | .3154 |
| 80 | .1631 | .1811 | .1977 | .2131 | .2276 | .2413 | .2542 | .2666 | .2784 | .2897 | .3005 |
| 90 | .1470 | .1636 | .1788 | .1931 | .2065 | .2192 | .2313 | .2428 | .2538 | .2645 | .2747 |
| 100 | .1339 | .1492 | .1633 | .1765 | .1889 | .2008 | .2121 | .2229 | .2333 | .2432 | .2529 |
| 110 | .1229 | .1371 | .1502 | .1625 | .1741 | .1852 | .1958 | .2060 | .2158 | .2252 | .2343 |
| 120 | .1136 | .1268 | .1390 | .1506 | .1615 | .1719 | .1819 | .1915 | .2007 | .2096 | .2182 |
| 130 | .1056 | .1179 | .1294 | .1403 | .1506 | .1604 | .1698 | .1788 | .1876 | .1960 | .2042 |
| 140 | .0986 | .1102 | .1211 | .1313 | .1410 | .1503 | .1592 | .1678 | .1761 | .1841 | .1919 |
| 150 | .0925 | .1035 | .1137 | .1234 | .1326 | .1414 | .1498 | .1580 | .1659 | .1735 | .1810 |
| 200 | .0706 | .0792 | .0872 | .0948 | .1021 | .1091 | .1158 | .1223 | .1287 | .1348 | .1408 |
| 300 | .0480 | .0539 | .0595 | .0648 | .0699 | .0749 | .0796 | .0843 | .0888 | .0932 | .0975 |
| 400 | .0363 | .0409 | .0451 | .0492 | .0532 | .0570 | .0607 | .0643 | .0678 | .0712 | .0746 |
| 500 | .0292 | .0329 | .0364 | .0397 | .0429 | .0460 | .0490 | .0519 | .0548 | .0576 | .0604 |
| 600 | .0244 | .0275 | .0305 | .0332 | .0359 | .0386 | .0411 | .0436 | .0460 | .0484 | .0507 |
| 700 | .0210 | .0237 | .0262 | .0286 | .0309 | .0332 | .0354 | .0375 | .0396 | .0417 | .0437 |
| 800 | .0184 | .0208 | .0230 | .0251 | .0271 | .0291 | .0311 | .0330 | .0348 | .0367 | .0384 |
| 900 | .0164 | .0185 | .0205 | .0224 | .0242 | .0260 | .0277 | .0294 | .0311 | .0327 | .0343 |
| 1000 | .0148 | .0167 | .0184 | .0202 | .0218 | .0234 | .0250 | .0265 | .0280 | .0295 | .0309 |

**Upper percentage points of 0.900 of theta(p,m,n),
the largest eigenvalue of |B-theta(W+B)|=0,when s=8**

m

| n | 0 | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 |
|---|---|---|---|---|---|---|---|---|---|---|
| 5 | 0.8517 | 0.8702 | 0.8845 | 0.8959 | 0.9052 | 0.9124 | 0.9225 | 0.9563 | 0.9570 | 0.9882 |
| 6 | 0.8183 | 0.8396 | 0.8563 | 0.8698 | 0.8808 | 0.8897 | 0.8978 | 0.9014 | 0.9096 | 0.9727 |
| 7 | 0.7862 | 0.8099 | 0.8286 | 0.8439 | 0.8566 | 0.8667 | 0.8781 | 0.8788 | 0.9019 | 0.9532 |
| 8 | 0.7558 | 0.7814 | 0.8019 | 0.8187 | 0.8328 | 0.8446 | 0.8559 | 0.8763 | 0.8994 | 0.9279 |
| 9 | 0.7270 | 0.7542 | 0.7761 | 0.7943 | 0.8096 | 0.8226 | 0.8347 | 0.8428 | 0.8664 | 0.8882 |
| 10 | 0.7000 | 0.7283 | 0.7515 | 0.7708 | 0.7872 | 0.8012 | 0.8139 | 0.8237 | 0.8385 | 0.8695 |
| 11 | 0.6744 | 0.7038 | 0.7280 | 0.7482 | 0.7656 | 0.7806 | 0.7937 | 0.8070 | 0.8163 | 0.8390 |
| 12 | 0.6505 | 0.6807 | 0.7056 | 0.7267 | 0.7448 | 0.7606 | 0.7746 | 0.7874 | 0.8013 | 0.8057 |
| 13 | 0.6280 | 0.6587 | 0.6843 | 0.7061 | 0.7248 | 0.7412 | 0.7555 | 0.7688 | 0.7826 | 0.7865 |
| 14 | 0.6069 | 0.6381 | 0.6641 | 0.6864 | 0.7057 | 0.7227 | 0.7376 | 0.7513 | 0.7665 | 0.7798 |
| 15 | 0.5870 | 0.6185 | 0.6449 | 0.6676 | 0.6874 | 0.7048 | 0.7202 | 0.7351 | 0.7468 | 0.7731 |
| 16 | 0.5683 | 0.5999 | 0.6267 | 0.6498 | 0.6699 | 0.6877 | 0.7036 | 0.7182 | 0.7341 | 0.7421 |
| 17 | 0.5507 | 0.5824 | 0.6094 | 0.6327 | 0.6532 | 0.6713 | 0.6874 | 0.7028 | 0.7153 | 0.7347 |
| 18 | 0.5340 | 0.5658 | 0.5929 | 0.6164 | 0.6371 | 0.6555 | 0.6720 | 0.6866 | 0.6981 | 0.7066 |
| 19 | 0.5183 | 0.5500 | 0.5772 | 0.6009 | 0.6218 | 0.6404 | 0.6572 | 0.6719 | 0.6841 | 0.6935 |
| 20 | 0.5035 | 0.5351 | 0.5623 | 0.5861 | 0.6071 | 0.6259 | 0.6428 | 0.6579 | 0.6713 | 0.6799 |
| 22 | 0.4761 | 0.5074 | 0.5345 | 0.5583 | 0.5796 | 0.5986 | 0.6159 | 0.6313 | 0.6450 | 0.6560 |
| 24 | 0.4514 | 0.4823 | 0.5092 | 0.5330 | 0.5542 | 0.5734 | 0.5908 | 0.6067 | 0.6205 | 0.6319 |
| 26 | 0.4291 | 0.4595 | 0.4861 | 0.5097 | 0.5309 | 0.5501 | 0.5676 | 0.5837 | 0.5982 | 0.6101 |
| 28 | 0.4089 | 0.4387 | 0.4649 | 0.4883 | 0.5094 | 0.5286 | 0.5461 | 0.5622 | 0.5768 | 0.5897 |
| 30 | 0.3904 | 0.4196 | 0.4454 | 0.4686 | 0.4895 | 0.5085 | 0.5260 | 0.5422 | 0.5569 | 0.5699 |
| 35 | 0.3507 | 0.3784 | 0.4031 | 0.4254 | 0.4457 | 0.4643 | 0.4816 | 0.4975 | 0.5123 | 0.5258 |
| 40 | 0.3182 | 0.3444 | 0.3679 | 0.3893 | 0.4089 | 0.4270 | 0.4438 | 0.4595 | 0.4741 | 0.4874 |
| 45 | 0.2912 | 0.3160 | 0.3383 | 0.3588 | 0.3776 | 0.3951 | 0.4114 | 0.4267 | 0.4410 | 0.4542 |
| 50 | 0.2684 | 0.2918 | 0.3131 | 0.3327 | 0.3507 | 0.3676 | 0.3833 | 0.3982 | 0.4121 | 0.4250 |
| 55 | 0.2488 | 0.2711 | 0.2913 | 0.3100 | 0.3274 | 0.3436 | 0.3588 | 0.3731 | 0.3866 | 0.3994 |
| 60 | 0.2319 | 0.2531 | 0.2724 | 0.2903 | 0.3069 | 0.3225 | 0.3371 | 0.3510 | 0.3641 | 0.3766 |
| 65 | 0.2172 | 0.2373 | 0.2557 | 0.2728 | 0.2888 | 0.3038 | 0.3179 | 0.3313 | 0.3441 | 0.3562 |
| 70 | 0.2042 | 0.2234 | 0.2410 | 0.2574 | 0.2727 | 0.2871 | 0.3008 | 0.3137 | 0.3261 | 0.3378 |
| 75 | 0.1926 | 0.2110 | 0.2278 | 0.2435 | 0.2583 | 0.2722 | 0.2854 | 0.2979 | 0.3098 | 0.3212 |
| 80 | 0.1824 | 0.1999 | 0.2160 | 0.2311 | 0.2453 | 0.2587 | 0.2715 | 0.2836 | 0.2952 | 0.3062 |
| 90 | 0.1647 | 0.1808 | 0.1958 | 0.2097 | 0.2229 | 0.2354 | 0.2473 | 0.2587 | 0.2696 | 0.2800 |
| 100 | 0.1502 | 0.1651 | 0.1790 | 0.1920 | 0.2043 | 0.2159 | 0.2271 | 0.2378 | 0.2480 | 0.2578 |
| 110 | 0.1380 | 0.1519 | 0.1648 | 0.1770 | 0.1885 | 0.1994 | 0.2099 | 0.2200 | 0.2296 | 0.2389 |
| 120 | 0.1277 | 0.1406 | 0.1527 | 0.1641 | 0.1749 | 0.1852 | 0.1951 | 0.2046 | 0.2138 | 0.2226 |
| 130 | 0.1188 | 0.1309 | 0.1423 | 0.1530 | 0.1632 | 0.1730 | 0.1823 | 0.1913 | 0.2000 | 0.2084 |
| 140 | 0.1110 | 0.1225 | 0.1332 | 0.1433 | 0.1530 | 0.1622 | 0.1711 | 0.1796 | 0.1879 | 0.1958 |
| 150 | 0.1042 | 0.1150 | 0.1252 | 0.1348 | 0.1439 | 0.1527 | 0.1611 | 0.1692 | 0.1771 | 0.1847 |
| 200 | 0.0798 | 0.0882 | 0.0962 | 0.1038 | 0.1111 | 0.1181 | 0.1248 | 0.1313 | 0.1377 | 0.1438 |
| 300 | 0.0543 | 0.0602 | 0.0658 | 0.0711 | 0.0762 | 0.0812 | 0.0860 | 0.0907 | 0.0952 | 0.0997 |
| 400 | 0.0412 | 0.0457 | 0.0500 | 0.0541 | 0.0580 | 0.0619 | 0.0656 | 0.0692 | 0.0728 | 0.0762 |
| 500 | 0.0331 | 0.0368 | 0.0403 | 0.0436 | 0.0469 | 0.0500 | 0.0530 | 0.0560 | 0.0589 | 0.0618 |
| 600 | 0.0277 | 0.0308 | 0.0338 | 0.0366 | 0.0393 | 0.0419 | 0.0445 | 0.0470 | 0.0495 | 0.0519 |
| 700 | 0.0238 | 0.0265 | 0.0290 | 0.0315 | 0.0338 | 0.0361 | 0.0383 | 0.0405 | 0.0426 | 0.0447 |
| 800 | 0.0209 | 0.0232 | 0.0255 | 0.0276 | 0.0297 | 0.0317 | 0.0337 | 0.0356 | 0.0375 | 0.0393 |
| 900 | 0.0186 | 0.0207 | 0.0227 | 0.0246 | 0.0265 | 0.0283 | 0.0300 | 0.0317 | 0.0334 | 0.0351 |
| 1000 | 0.0168 | 0.0187 | 0.0205 | 0.0222 | 0.0239 | 0.0255 | 0.0271 | 0.0286 | 0.0301 | 0.0316 |